

Confidence Intervals:

1. C.I. width: confidence $\uparrow$, width $\uparrow$;
sample size $\uparrow$, width $\downarrow$;
population "noise" $\uparrow$, width $\uparrow$
2. C.I. width = $(2)(\text{Margin of Error})$

Interval:

1. General: $\text{point estimator} \pm (\text{table value})(\text{STD error})$
2. Z-interval
 - $\bar{x} \pm Z\sigma_{\bar{x}} \quad \sigma_{\bar{x}} = \frac{\sigma}{\sqrt{n}}$
3. t-interval
 - $\bar{x} \pm ts_{\bar{x}} \quad s_{\bar{x}} = \frac{s}{\sqrt{n}}$

Test Statistic:

1. General: $\frac{\text{point estimator} - \text{null value}}{\text{STD error}}$
2. Z: $\frac{\bar{x} - \text{null value}}{\sigma_{\bar{x}}}$
3. t: $\frac{\bar{x} - \text{null value}}{s_{\bar{x}}}$

Sample Size for Desired Margin of Error:

1. $n = \left(\frac{Z(\sigma)}{E}\right)^2$ where E is the desired margin of error

Hypothesis Testing:

1. H_o = null hypothesis
2. H_a = alternative hypothesis
3. α = Type I Error Probability = $P(\text{Type I Error})$
4. β = Type II Error Probability = $P(\text{Type II Error})$
5. $\text{Power} = 1 - \beta$; $\text{Power} = P[\text{Reject } H_o]$
 - a. sample size $\uparrow$, power $\uparrow$
 - b. effect size $\uparrow$, power $\uparrow$
 - c. α -level $\uparrow$, power $\uparrow$
 - d. noise $\uparrow$, power $\downarrow$
 - e. as $n \uparrow$, power curve becomes tighter and closer to the ideal
6. reject H_o if p-value is below α -level

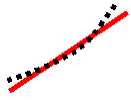
	Reject H_o	Fail to Reject H_o
H_o True	Type I Error	Correct
H_o False	Power	Type II Error

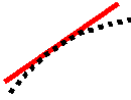
****Make a conclusion in terms of/in the context of the problem****

t-Distribution (σ unknown):

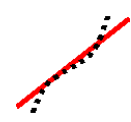
1. more spread than STD normal; symmetric at 0; as the degrees of freedom $\uparrow$, becomes more like STD Normal
2. degrees of freedom = $n - 1$

Normal Probability Plot:

 Right Skew - If the plotted points appear to bend up and to the left of the normal line that indicates a long tail to the right.

 Left Skew - If the plotted points bend down and to the right of the normal line that indicates a long tail to the left.

 Short Tails - An S shaped-curve indicates shorter than normal tails, i.e. less variance than expected.

 Long Tails - A curve which starts below the normal line, bends to follow it, and ends above it indicates long tails. That is, you are seeing more variance than you would expect in a normal distribution.

1. $x = \mu + Z\sigma$
2. Ordered data $\rightarrow$ normal score (Z)

Sign Test:

1. Nonparametric – use median in hypothesis
2. $x_i - \text{null value} = \#_i \quad n = i$
3. If $x_i = \text{null value}$ then drop that value and adjust n.
4. $B = \#$ of positive signs (test statistic)
5. Binomial ($n, p=0.05$) $\Rightarrow$ Reference distribution
 $\hookrightarrow H_o$ true
6. Binomial Distribution (n, p)
 - a. $n = \#$ of trials; $p = P[\text{success}]$
 - b. trials must be Independent of one another
 - c. $p(x) = \begin{cases} \binom{n}{x} p^x (1-p)^{n-x} & x = 0, 1, \dots, n \\ 0 & \text{Otherwise} \end{cases}$
 - d. $E(X) = np$
 - e. $V(X) = np(1-p)$

Comparing Means-Paired Data:

1. Normality: t-procedures
2. Nonparametric: Sign Test
 - a. Dependent samples
 - b. mean difference: $\mu_d = \mu_1 - \mu_2$
 - c. sample size $\rightarrow n = \#$ of pairs
 - d. pt. estimator: $\bar{d}$
 - e. STD error: $s_{\bar{d}} = \frac{s_e}{\sqrt{n}}$
 - f. C.I.: $\bar{d} \pm t \cdot s_{\bar{d}}$
 - g. T.S.: $t = \frac{\bar{d} - \text{null value}}{s_{\bar{d}}}$

3. Nonparametric: Signed-Rank Test
 - a. Calculate the differences; delete any zeros
 - b. absolute value each difference
 - c. assign ranks; ties get the mean rank
 - d. create sign ranks (bring back signs)
 - e. total negative ranks: T^-
total positive ranks: T^+
 - under H_0 true: $T^- = T^+ = \frac{n(n+1)}{4}$

Comparing Means-Independent Data:

1. Normal: Two-sample t-inference
 - a. parameter of interest: $\mu_1 - \mu_2$
 - b. pt. estimate: $\bar{Y}_1 - \bar{Y}_2$

$$E(\bar{Y}_1 - \bar{Y}_2) = \mu_1 - \mu_2$$

$$Var(\bar{Y}_1 - \bar{Y}_2) = \frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}$$
 - c. STD error: $s_{\bar{Y}_1 - \bar{Y}_2} = \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$
 - d. T.S.: $t = \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{s_{\bar{Y}_1 - \bar{Y}_2}}$
 - e. C.I.: $(\bar{Y}_1 - \bar{Y}_2) \pm t \cdot s_{\bar{Y}_1 - \bar{Y}_2}$
2. Nonparametric: Sum Rank Test
 - a. Arrange the data for each group in increasing order
 - b. rank the observations as though they were all drawn from the same (assign tied data the average of the ranks)
 - c. total the ranks for the 2 groups T_1, T_2
 - d. check $T_1 + T_2 = \frac{(n_1 + n_2)(n_1 + n_2 + 1)}{2}$
 - e. use T_1 as the test statistic
 - f. make conclusion with table on page 1176

Correlation:

- one quantitative response-dependent (y)
- one quantitative predictor-independent/control (x)
- Association/Relationship/Co-related/Co-vary
- Scatter plot (must have correlation)

Questions of Interests:

- Strength (tightness of trend)
- Type (linear vs. nonlinear)
- Direction (positive/negative association)

Linear Correlation Coefficient: $\frac{obs - mean}{stdev}$

$$r = \frac{\sum(x - \bar{x})(y - \bar{y})}{\sqrt{\sum(x - \bar{x})^2 \sum(y - \bar{y})^2}} = \frac{s_{xy}}{(n-1) s_x s_y}$$

population correlation: ρ (parameter)

sample correlation: r (statistic)

Linear Regression:

modeling relationship(s) among variables

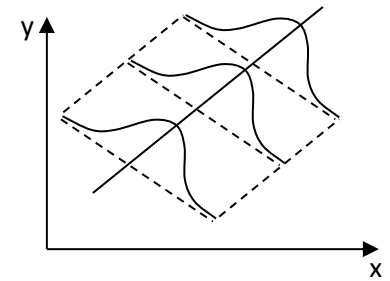
modeling mean/location of Y as a linear function of X

$$E(X|Y) = \beta_0 + \beta_1 X$$

$$Y = \beta_0 + \beta_1 X + \varepsilon$$

3 parameters of interest:

1. intercept - β_0
2. slope - β_1
3. variance - σ^2



Sums of Squares:

$$\sum (observed - predicted)^2 = \sum (errors)^2$$

$$\sum (y - \bar{y})^2 = \sum (errors)^2$$

1. $\sum (y - \bar{y})^2 = \sum (y - \hat{y})^2 + \sum (\hat{y} - \bar{y})^2$
 $SS_{total} = SS_{error} + SS_{regression}$
- d.f. - $n - 1 = \left(n - \frac{\#(beta)}{\text{parameters}} \right) + \# \text{predictors}$
2. Best Line
 - a. minimizes the sum of squared errors
 - b. $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 X$
3. Estimates:
 - a. $\hat{\beta}_1 = \frac{s_{xy}}{s_{xx}} = r \frac{s_y}{s_x}$
 - b. $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$

Mean Squares Error:

1. General: $MS = \frac{SS}{df}$
2. Regression: $MS_{reg} = \frac{SS_{reg}}{d.f. \cdot reg}$
3. Error: $MS_{err} = \sigma^2 = \frac{\sum (y - \bar{y})^2}{n - \frac{\# of (beta)}{\text{parameters}}} = \frac{SS_{err}}{d.f. \cdot err}$
4. $R^2 = \frac{\text{explained}}{\text{unexplained}} = \frac{SS_{regression}}{SS_{error}} \quad 0 \leq R^2 \leq 1$
5. $SE_{\hat{\beta}_0} = \sigma \sqrt{\frac{1}{n} + \frac{\bar{x}^2}{s_{xx}}} = \sqrt{MS_E} \cdot \sqrt{\frac{1}{n} + \frac{\bar{x}^2}{s_{xx}}}$
6. $SE_{\hat{\beta}_1} = \frac{\sigma}{\sqrt{s_{xx}}} = \frac{\sqrt{MS_E}}{\sqrt{s_{xx}}}$
 Note: to reduce $SE_{\hat{\beta}_1}$, increase s_{xx}
7. C.I. for $\hat{\beta}_0$, intercept: $\hat{\beta}_0 \pm t \cdot SE_{\hat{\beta}_0}$
 $d.f. = n - \frac{\# of (beta)}{\text{parameters}}$
8. C.I. for $\hat{\beta}_1$, slope: $\hat{\beta}_1 \pm t \cdot SE_{\hat{\beta}_1}$
 $d.f. = n - \frac{\# of (beta)}{\text{parameters}}$
9. $H_0: \beta_1 = \frac{null}{value} \rightarrow \text{typically } 0 \leftrightarrow \text{no relationship}$
 $H_a: \beta_1 <, \neq, > \frac{null}{value}$
10. T.S. - $t = \frac{\text{estimate} - \frac{null}{value}}{SE_{\hat{\beta}_1}}$

Analysis of Variance (ANOVA):

Source	d.f.	SS	MS	F	p-value
Regression	# predictors	SS_{reg}	MS_{reg}	$\frac{MS_{reg}}{MS_{err}}$	Table Value
Residual Error	$n - \#(\text{beta})$ parameters	SS_{err}	MS_{err}	-----	-----
Total	$n - 1$	SS_{total}	-----	-----	-----

**error = residual = $y - \bar{y}$

Model Utility Test (F-test):

- H_0 : not useful model $y = \mu + \varepsilon$
 H_a : model useful predictor coefficient $\neq 0$
- F-dist Δ (Table 8; pg 1181-1193)
 - numdf = $df_1 = df_{reg}$
 - dendf = $df_2 = df_{res}$
- $SE_{\hat{y}} = \sqrt{MS_E} \cdot \sqrt{\frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}}$
- C.I. for the mean of $Y|X$: $\hat{y} \pm t \cdot SE_{\hat{y}}$
-use df_{res}
- P.I. for $Y|X$: $\hat{y} \pm t \cdot \sqrt{MS_E} \cdot \sqrt{1 + \frac{1}{n} + \frac{(x - \bar{x})^2}{S_{xx}}}$
-use df_{res}
- Watch for:
 - Large residuals
 - non-random pattern
 - unusual x-values/predictor values

Lack of Fit:

- $SS_{residues} = SS_{error} + SS_{lack \text{ of fit}}$
- d.f. - $n - \frac{\# \text{ of } (\text{beta})}{\text{parameters}} = (n - m) + \left(m - \frac{\# \text{ of } (\text{beta})}{\text{parameters}}\right)$
 $m = \# \text{ of distinct } x \text{ values}$
- H_0 : no lack of fit; appropriate model
 H_a : lack of fit; poor model
- T.S. - $F = \frac{MS_{LOF}}{MS_{PE}} \quad df_1 = df_{LOF}; df_2 = df_{PE}$

Multiple Linear Regression:

- $y = \text{location} + \text{spread} = \beta_0 + \beta_1 X_1 + \dots + \beta_k X_k + \varepsilon$
 k : predictors; $k + 1$: β parameters
- Estimate of $\sigma^2 = MS_E = \frac{SS_E}{n - \frac{\# \text{ of } (\text{beta})}{\text{parameters}}}$
- Estimate of $\sigma = \sqrt{MS_E}$
- $R^2 = \frac{SS_{reg}}{SS_{total}} = 1 - \frac{SS_{err}}{SS_{tot}}$
Note: can increase R^2 by adding polynomial terms
- $R_{adj}^2 = 1 - \frac{\frac{SS_E}{df_E}}{\frac{SS_{tot}}{df_{tot}}}$
- H_0 : all predictor coefficients = 0; not useful model
 H_a : at least one predictor coefficient $\neq 0$; useful model

$$7. \text{ T.S. } - F = \frac{MS_{reg}}{MS_{err}} = \frac{\frac{R^2}{k}}{\frac{1-R^2}{n-(k+1)}}$$

$$df_1 = df_{reg} = k; df_2 = df_{err} = n - (k + 1)$$

- C.I. for β_i : $\hat{\beta}_i \pm t \cdot SE_{\hat{\beta}_i}$
-use df_{res}

- T.S. for β_i : $\frac{\hat{\beta}_i - \text{null value}}{SE_{\hat{\beta}_i}}$
-use df_{res}

Testing a Subset of Parameters:

- Full Model vs. Reduced Model
- H_0 : insignificant; $\beta_i = 0$
 H_a : at least one $\beta_i \neq 0$
- Estimation of the mean response at predictor values:
 - C.I.: $\hat{y} \pm t \cdot SE_{\hat{y}}$ -use df_{res}
- Prediction of the response at predictor values (wider):
 - P.I.: $\hat{y} \pm t \cdot \sqrt{MS_E + (SE_{\hat{y}})^2}$ -use df_{res}
- Partial F-Test: $F = \frac{\frac{SS_{reg,2} - SS_{reg,1}}{df_{reg,2} - df_{reg,1}}}{\frac{SS_{reg,1}}{df_{reg,1}}}$
where the model for 1 is larger than 2

Matrix Arithmetic:

- $\underline{Y} = X\underline{\beta} + \underline{\varepsilon}$
 $\underline{Y} = n \times 1; X = n \times (k + 1); \underline{\beta} = (k + 1) \times 1;$
 $\underline{\varepsilon} = n \times 1$
- Least Squares Estimation: $\underline{\hat{\beta}} = (X'X)^{-1}X'\underline{Y}$
 $\underline{\hat{\beta}} = (k + 1) \times 1; X' = (k + 1) \times n;$
 $X = n \times (k + 1); \underline{Y} = n \times 1$
Note: each estimate of the coefficient is a linear combination of the y 's
- $X\underline{\hat{\beta}} = \underline{\hat{Y}}$
 $\underline{\hat{Y}} = n \times 1$
- Hat Matrix: $[X(X'X)^{-1}X']\underline{Y} = H\underline{Y}$
- Residual: $\underline{Y} - \underline{\hat{Y}} = [I\underline{Y} - H\underline{Y}] = [I - H]\underline{Y}$
- Estimated: $Var(\underline{\hat{\beta}}) = MS_E[(X'X)^{-1}]$
Small?
-determinate small
-trace small
-maximum eigenvalue small

Things that could go wrong:

1. Wrong Model: nonlinear
2. non-constant variance
3. outliers

How to diagnose?:

1. graphical analysis
 - a. why not just plot the y's (not the same mean, not the same distribution) => go to the residuals

Residuals:

1. Standardized
 - a. $\frac{e_i}{\sigma\sqrt{1-h_{ii}}}$
 - b. significant: ± 2
2. "Studentized"
 - a. Internal – standardized (Minitab); student (SAS)
 - i. $\frac{e_i}{\sqrt{MS_{E(i)}\sqrt{1-h_{ii}}}}$
 - b. External – deleted t residual (Minitab); rstudent (SAS)
 - i. $\frac{e_i}{\sqrt{MS_{E(i)}\sqrt{1-h_{ii}}}}$
 - ii. $MS_{E(i)}$: mean square error fitted with the i^{th} case deleted
 - c. Plots
 - i. histogram, boxplot, npp, ...
 - ii. scatterplots (vs): time, $\hat{y}$, x
3. Something wrong?
 - a. change model: add predictors; go to non-linear model
 - b. go away from least squares estimation: use a different method of model fitting
 - c. if errors are normal with constant variance, consider transformation of x
 - d. if errors are non-normal and/or constant variance, consider transformation of y

Potential/Leverage:

1. focus on the predictors; measure of how far the predictor values are away from the center of the x's
 - a. General
 - i. the h_{ii} can be interpreted as the distance the i^{th} case of predictors is from the center of all the predictors
 - 1) large values, potential/leverage to influence
 - 2) guidelines for "too big": $\frac{2(\# \text{ predictors})}{n}$

Influence:

1. Consideration of both the x (predictors) and y (response)
2. Methods of Comparison:
 - a. DFBetas (SAS only)
 - i. influence of the i^{th} case on estimating β_k
 - ii. $\frac{\hat{\beta}_k - \hat{\beta}_{k(i)}}{\sqrt{MS_{E(i)} \cdot c_{kk}}}$
 c_{kk} is the k^{th} diagonal element of $(X'X)^{-1}$
 - iii. Guidelines:
 - 1) > 1 , further investigation
 - 2) $> \frac{2}{\sqrt{n}}$, large data set; $n \geq 100$
 - b. DFFits
 - i. measure the influence the i^{th} obs has on the fitted value $\hat{y}_i$
 - ii. $DFFits_i = \frac{\hat{y}_i - \hat{y}_{k(i)}}{\sqrt{MS_{E(i)} \cdot h_{ii}}}$
 - iii. Guidelines:
 - 1) $|DFFits| > 1$, small/medium data sets
 - 2) $|DFFits| > 2\sqrt{\frac{\# \text{ parameters}}{n}}$, $n \geq 100$
 - c. Cooks Distance
 - i. consider the influence of the i^{th} case on all fitted values "aggregate influence measure"
 - ii. $p = \# \text{ parameters}$
$$D_i = \frac{\sum_{j=1}^n (\hat{y}_j - \hat{y}_{j(i)})^2}{p \cdot MS_E}$$
 - iii. Guidelines:
 - 1) $D_i > 1$, unreasonable
 - d. Cov Ratio
 - i. $\frac{MS_{E(i)}(X_{(i)}^i X_{(i)})^{-1}}{|MS_E(X'X)^{-1}|}$
 - ii. Guidelines:
 - 1) $|CovRatio - 1| \geq \frac{3p}{n}$

Model Building:

1. plot y vs. each X (matrix scatterplot, correlation matrix)
2. multicollinearity: relationship amongst predictors (independent variables)
 - a. effects
 - i. insignificant t-tests when you expect significance
 - ii. signs for the betas opposite of what you expect
 - iii. error significant in the computation; numerical analysis

$$b. SE_{\hat{\beta}_i} = \sqrt{MS_E \left(\frac{1}{1-R_i^2} \right)}$$

↳ variance of inflation

Guidelines: $VIF > 10$

where R_i^2 is the R^2 when x_i is regressed on the other predictors

3. Compare Criteria
 - a. R^2 : high
 - b. MS_E : low
 - c. $R^2 - adj$: high
 - d. $PRESS = \sum_{i=1}^n (y_i - \hat{y}_{(i)})^2$: low
↳ predicted residuals of SS
 - e. Mallows' $C_p \approx p$
 - i. $C_p = \frac{SS_{E_p}}{MS_E} - (n - 2p)$
4. Model Selection:
 - a. Backward Selection
 - i. start big-remove predictors: fit the model with all predictors; pick out biggest p-value; refit model => continue until all predictors are significant
 - b. Forward Selection
 - i. start small-add predictors: consider all SLR models; add the predictor with the smallest p-value => continue until no longer significant
 - c. Stepwise Selection
 - i. combination of backward and forward selection: check overall model to make sure that all predictors are still useful; if not, it is removed from the model

One-Way ANOVA:

$$y_{ij} = \mu_i + \varepsilon_{ij} = \mu + \tau_i + \varepsilon_{ij}$$

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k$$

$$H_0: \tau_1 = \tau_2 = \dots = \tau_k = 0$$

1. compare the means of three or more groups
2. Hypotheses:

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k$$

$$H_a: \text{means are not all equal}$$

3. Requirements:
 - a. the k populations all have normal distributions
 - b. the k populations all have the same standard deviations $\sigma_1 = \sigma_2 = \dots = \sigma_k$
4. Notation:
 - a. y_{ij} = observation for the jth unit in the ith group
 - b. n_i = number of units in the ith group
 - c. n = total number of units in the study
 - d. $\bar{y}_{i.}$ = sample mean for the ith group
 - e. $\bar{y}_{..}$ = sample mean for the entire set of data (grand mean)
 - f. k = number of groups
5. Assessment of the within-group variation [error/noise]
 - a. $SS_{within} = (n_1 - 1)(s_1^2) + (n_2 - 1)(s_2^2) + \dots + (n_k - 1)(s_k^2)$
 - b. $df_{within} = n - k$
 - c. $MS_{within} = \frac{SS_{within}}{df_{within}} = \frac{(n_1-1)(s_1^2) + (n_2-1)(s_2^2) + \dots + (n_k-1)(s_k^2)}{n_1 + n_2 + \dots + n_k - k}$
6. Assessment of the between-group variation [treatment]
 - a. $SS_{between} = \sum n_i (\bar{y}_{i.} - \bar{y}_{..})^2$
 $SS_{between} = n \sum (\bar{y}_{i.} - \bar{y}_{..})^2$, if all groups have the same number of units
 - b. $df_{between} = k - 1$
 - c. $MS_{between} = \frac{SS_{between}}{df_{between}}$
7. $\bar{y}_{..} = \frac{n_1 \bar{y}_{1.} + n_2 \bar{y}_{2.} + \dots + n_k \bar{y}_{k.}}{n_1 + n_2 + \dots + n_k}$
8. F-test: $F = \frac{MS_{between}}{MS_{within}}$
9. $SS_{total} = \sum \sum (y_{ij} - \bar{y}_{..})^2$
10. $s = \sqrt{\frac{SS_{total}}{n-1}}$
11. $R^2 = \frac{SS_{between}}{SS_{total}}$

Multiple Comparisons:

α_I = individual type I error rate

α_E = experiment wise type I error rate

1. Fisher's LSD (liberal)

$$a. t_{ij} = \frac{(\bar{y}_{i.} - \bar{y}_{j.})}{SE_{\bar{y}_{i.} - \bar{y}_{j.}}}$$

$$\text{where } SE_{\bar{y}_{i.} - \bar{y}_{j.}} = \sqrt{MS_{within} \left(\frac{1}{n_i} + \frac{1}{n_j} \right)}$$

$$b. df = df_{within}$$

$$c. CI: \bar{y}_{i.} - \bar{y}_{j.} \pm t \cdot SE_{\bar{y}_{i.} - \bar{y}_{j.}}$$

2. Kruskal-Wallis

a. assign ranks to all the data, averaging for repeated values

b. determine the mean rank within each group $\bar{R}_i$

c. If H_0 true, the $\bar{R}_i$ "equal"

$$d. H \text{ test stat: } \frac{12}{n(n+1)} \sum n_i (\bar{R}_i - \bar{R})^2$$

reject large H (pg 428)

$$e. \text{ total ranks} = \frac{n(n+1)}{2}$$

$$f. \text{ mean rank} = \frac{n+1}{2}$$

3. Bonferroni Correction (conservative)

a. want $\alpha_E = \alpha \Rightarrow$ choose $\alpha_I = \frac{\alpha}{m}$ for m independent tests

b. LSD - $t \cdot SE_{\bar{y}_{i.} - \bar{y}_{j.}}$
find t with $\frac{\alpha_I}{2}$ in each tail (t-tables)

$$c. CI: \bar{y}_{i.} - \bar{y}_{j.} \pm t \cdot SE_{\bar{y}_{i.} - \bar{y}_{j.}}$$

4. Tukey's W Procedure (HSD) (mod-conservative)

a. Table 10

$$b. \text{ common sample size: } w = q_{\alpha}(\#trt, df_{err}) \sqrt{\frac{MS_E}{n}}$$

$$c. \text{ uncommon sample size: } w = \frac{q_{\alpha_E}(\#trt, df_{err})}{\sqrt{2}} \sqrt{MS_E \left(\frac{1}{n_i} + \frac{1}{n_j} \right)}$$

$$d. |\bar{y}_{i.} - \bar{y}_{j.}| \geq w$$

$$e. CI: \bar{y}_{i.} - \bar{y}_{j.} \pm w$$

5. SNK (mod-liberal)

$$a. SNK = q_{\alpha_E}(\# \text{ means}, df_{err}) \sqrt{\frac{MS_E}{n}}$$

$$b. |\bar{y}_{i.} - \bar{y}_{j.}| \geq SNK$$

6. Dunnett's Compare to Control

a. Are treatments different than the control?

b. two sided test: $|\bar{y}_{i.} - \bar{y}_{cont}| \geq$

$$d_{\alpha_E}(\# \text{ non control}, df_{err}) \sqrt{\frac{2MS_E}{n}}$$

Blocking (Randomized Complete Block Design):

$$y_{ij} = \mu_{ij} + \varepsilon_{ij} = \mu + \tau_i + \beta_j + \varepsilon_{ij}$$

$$\sum \tau_t = 0 \quad \sum \beta_b = 0$$

total # obs = bt (with no repeats)

$$SS_{total} = SS_{block} + SS_{trt}$$

RCBD ANOVA Table:

Source	d.f.	SS	MS	F	p-value
Treatment	t-1	SS_{trt}	MS_{trt}	$\frac{MS_{trt}}{MS_{err}}$	Table Value
Block	b-1	SS_{blk}	MS_{blk}	$\frac{MS_{blk}}{MS_{err}}$	Table Value
Error	(b-1)(t-1)	SS_{err}	MS_{err}	-----	-----
Total	bt-1	SS_{totl}	-----	-----	-----

1. Treatment Effective?

$$H_0: \tau_1 = \tau_2 = \dots = \tau_k = 0$$

$$H_a: \text{not all } \tau = 0$$

2. Blocking Effective?

$$H_0: \beta_1 = \beta_2 = \dots = \beta_k = 0$$

$$H_a: \text{not all } \beta = 0$$

3. Comparison Tests:

a. Tukey's Pairwise comparisons for RCB

$$i. |\bar{y}_{i.} - \bar{y}_{j*}| \geq q_{\alpha}(\#trt, df_{err}) \sqrt{\frac{MS_E}{b}}$$

b. Bonferroni

$$i. |\bar{y}_{i.} - \bar{y}_{j*}| \geq t \sqrt{MS_E \left(\frac{1}{b} + \frac{1}{b} \right)}$$

Factorial Structure:

Factor A – a levels

Factor B – b levels

Balanced Design:

ab = # treatments

abn = # observations (equal replication)

$$y_{ijk} = \mu + \alpha_i + \beta_j + \alpha\beta_{ij} + \varepsilon_{ijk}$$

ME ME INT

$$SS_{tot} = SS_E + SS_A + SS_B + SS_{AB}$$

ANOVA:

Source	d.f.	SS	MS	F	p-value
A	a-1	SS_A	MS_A	$\frac{MS_A}{MS_{err}}$	Table Value
B	b-1	SS_B	MS_B	$\frac{MS_B}{MS_{err}}$	Table Value
AB	(a-1)(b-1)	SS_{AB}	MS_{AB}	$\frac{MS_{AB}}{MS_{err}}$	Table Value
Error	ab(n-1)	SS_{err}	MS_{err}	-----	-----
Total	abn-1	SS_{total}	-----	-----	-----

1. Test: $H_0: \alpha\beta_{ij} = 0 \rightarrow$ no interaction

a. No: Main Effects $\rightarrow$ no main effects

$$i. H_0: \alpha_i = 0$$

$$ii. H_0: \beta_j = 0$$

b. Yes: M/C between trt means

2. Simple Effects: comparison in which all but one factor is fixed at a certain level

$$a. \# \text{ of simple effects} = \binom{a}{2}b + \binom{b}{2}a$$

b. use M/C defined in Blocking