- 1)
- a. Latin Square Design with blocking variables Farm and Fertility. The treatment is the five types of fertilizers.
- b. There is significant evidence (p-value < 0.0001) the mean yields are different for the five fertilizers.

proc glm;

class Farm Fertility Fertilizers;
model Yield = Farm Fertility Fertilizers;
lsmeans Fertilizers / pdiff cl adj=tukey;
run;

The GLM Procedure

Class Level Information

Class	Leve	ls	Vá	alı	ıes	3		
Farm		5	1	2	3	4	5	
Fertility	5	1	2	3	4	5		
Fertilize	ers	5	Α	В	С	D	Ε	
	Observations Observations							25 25

The GLM Procedure

Dependent Variable: Yield

Source		DF	Sum Squa	of	Mean Square	F Value	Pr > F
Model		12	46.06720	000	3.83893333	9.88	0.0002
Error		12	12 4.663200		0.38860000		
Corrected To	tal	24	50.73040	000			
	R-Square	Coef	f Var	Root	MSE Yield	Mean	
	0.908079	8.7	45481	0.62	3378 7.12	8000	
Source		DF	Type I	SS	Mean Square	F Value	Pr > F
Farm Fertility		4	6.52240 11.26640	000	1.63060000 2.81660000	4.20 7.25	0.0236
Fertilizers		4	28.27840	000	7.06960000	18.19	<.0001

Evaluation #1 SKETCH OF SOULTIONS

Source	DF	Type III SS	Mean Square	F Value	Pr > F	
Farm	4	6.52240000	1.63060000	4.20	0.0236	
Fertility	4	11.26640000	2.81660000	7.25	0.0033	
Fertilizers	4	28.27840000	7.06960000	18.19	<.0001	

The GLM Procedure
Least Squares Means
Adjustment for Multiple Comparisons: Tukey

Fertilizers	Yield LSMEAN	LSMEAN Number
А	5.32000000	1
В	6.5600000	2
С	7.6400000	3
D	7.8800000	4
E	8.24000000	5

Least Squares Means for effect Fertilizers
Pr > |t| for H0: LSMean(i)=LSMean(j)

Dependent Variable: Yield

i/j	1	2	3	4	5
1 2 3 4 5	0.0537 0.0006 0.0002 <.0001	0.0537 0.1056 0.0380 0.0080	0.0006 0.1056 0.9710 0.5687	0.0002 0.0380 0.9710 0.8865	<.0001 0.0080 0.5687 0.8865
	Fertilizers	Yield LSMEAN	95% Confide	nce Limits	
	A B C D	5.320000 6.560000 7.640000 7.880000 8.240000	4.712584 5.952584 7.032584 7.272584 7.632584	5.927416 7.167416 8.247416 8.487416 8.847416	

Least Squares Means for Effect Fertilizers

		Difference Between	Simultane Confidence L	
i	j	Means	LSMean(i)-L	SMean(j)
1	2	-1.240000	-2.496643	0.016643
1	3	-2.320000	-3.576643	-1.063357
1	4	-2.560000	-3.816643	-1.303357
1	5	-2.920000	-4.176643	-1.663357
2	3	-1.080000	-2.336643	0.176643
2	4	-1.320000	-2.576643	-0.063357
2	5	-1.680000	-2.936643	-0.423357
3	4	-0.240000	-1.496643	1.016643
3	5	-0.600000	-1.856643	0.656643
4	5	-0.360000	-1.616643	0.896643

Using Tukey's W-procedure with $\alpha = 0.05$, $s_{\epsilon}^2 = MSE = 0.3886$, $q_{\alpha}(t, df_{Error}) = q_{\alpha}(5,12) = 4.52 \Rightarrow$

$$W = (4.52)\sqrt{\frac{0.3886}{5}} = 1.26 \Rightarrow$$

	Fertilizer									
	A	В	C	D	E					
Mean	5.32	6.56	7.64	7.88	8.24					
Grouping	a	ab	bc	c	c					

The following pairs of fertilizers have significantly different mean yields: (A,C), (A,D), (A,E), (B,D), (B,E)

Tukey controls experimentwise error rate see page 468 Chapter 9

15.26

A randomized complete block design with days as blocks and treatments consisting of the 3x4 temperature-pressure combinations. The twelve treatments would be randomly assigned to twelve samples on each of the three days so that one replication of the 3x4 factorial experiment would be observed each day. A diagram is given here:

		Day 1			Day 2		Day 3			
	T	emperatu	re	T	emperatu	re	Temperature			
Pressure	280°F	300°F	320°F	280°F	300°F	320°F	280°F	300°F	320°F	
100	S6	S1	S12	S9	S5	S1	S12	S3	S7	
150	S3	S11	S8	S6	S4	S10	S8	S9	S2	
200	S5	S 7	S4	S2	S3	S7	S4	S5	S11	
250	S10	S9	S2	S11	S8	S12	S6	S10	S1	

- 3) Note: a case can be made for Block Designs ... for example: considering grade level as a blocking factor ...
 - a. Because all the students were in the same grade, this is a completely randomized design with a 2x3 factorial treatment structure. Factor A-Sex and Factor B-Level of Abuse (3 levels). There are 30 reps of the complete experiment.
 - b. The grade level factor is considered as a third factor since age, as reflected by grade level, may interact with sex, because girls tend to mature more rapidly than boys. Thus, the design would be a completely randomized design with a 2x3x3 factorial treatment design: Factor A-Sex, Factor B-Level of Abuse (3 levels), Factor C-Grade Level (3 levels). There are 10 reps of the complete experiment.

4)

a. There are 12 treatments consisting of the 2 levels of Factor A, 3 levels of Factor B, and 2 levels of Factor C. In each block, randomly assign the numbers 1,2,...,12 to the 12 experimental units. The 12 treatments will then be randomly assigned to the 12 experimental units in each of the three blocks as seen in the following diagram:

		Blo	ck 1			Block 2			Block 3			Block 4				
	A	1	I	A 2	I	A1	A	12	A	1	P	12	A	1	A	12
Factor B	C1	C2	C1	C2	C1	C2	C1	C2	C1	C2	C1	C2	C1	C2	C1	C2
B1	U10	U8	U3	U12	U3	U11	U8	U4	U4	U9	U7	U3	U5	U3	U2	U4
B2	U7	U1	U6	U4	U6	U10	U2	U9	U10	U8	U5	U12	U8	U6	U7	U12
B3	U11	U9	U5	U2	U1	U7	U5	U12	U6	U1	U2	U11	U11	U1	U9	U10

b. The complete AOV table is given here:

implete 110 v tae	710 15 51 VCH 1	icic.			
Source	DF	SS	MS	F	p-value
Blocks	3	SST	SST/3	MST/MSE	
Factor A	1	SSA	SSA/1	MSA/MSE	
Factor B	2	SSB	SSB/2	MSB/MSE	
AB	2	SSAB	SSAB/2	MSAB/MSE	
Factor C	1	SSC	SSC/1	MSC/MSE	
AC	1	SSAC	SSAC/1	MSAC/MSE	
BC	2	SSBC	SSBC/2	MSBC/MSE	
ABC	2	SSABC	SSABC/2	MSABC/MSE	
Error	33	SSE	SSE/33		
Total	47	SSTot			

- a. Completely randomized design with a 3x2 factorial treatment structure and 10 reps.
- **b.** $y_{ijk} = \mu + \tau_i + \beta_j + (\tau \beta)_{ij} + \varepsilon_{ijk}$; i = 1, 2, 3; j = 1, 2; k = 1, ..., 10;

Where y_{ijk} is the attention span of the k^{th} child of Age i viewing Product j

 τ_i is the effect of the i^{th} Age on attention span

 β_j is the effect of the j^{th} Product on attention span

 $(\tau \beta)_{ij}$ is the interaction effect of the i^{th} Age and j^{th} Product on attention span

```
proc glm data=ad;
class age product;
model time = age|product / solution;
lsmeans age*product /out=abmeans;
output out=residata p=yhat rstudent=stdres;
run;

title2 "Profile/Interaction Plots";
symbol1 i=j;

proc gplot data=abmeans;
plot lsmean*age=product;
run;
```

The GLM Procedure

Class Level Information

Class	Levels	Values
age	3	A1 A2 A3
product	2	P1 P2

Number of Observations Read 60 Number of Observations Used 60

The GLM Procedure

Dependent Variable: time

			Sum	of				
Source		DF	Squa	-	Mean Squa	are F	Value	Pr > F
Model		5	4705.73	333	941.14	667	6.40	<.0001
Error		54	7944.00	000	147.11	111		
Corrected To	tal	59	12649.73	333				
	R-Square	Coef	f Var	Root	MSE t	ime Mean		
	0.372003	44.	48265	12.12	2894	27.26667		
Source		DF	Maron T	C C	Maan Cau	- E 1	770]0	D > E
source		DE	Type I	55	Mean Squa	are r	Value	Pr > F
age		2	1303.033		651.516		4.43	0.0166
product		1	2018.400		2018.400		13.72	0.0005
age*product		2	1384.300	000	692.150	300	4.70	0.0131
Source		DF	Type III	SS	Mean Squa	are F	Value	Pr > F
age		2	1303.033	333	651.516	667	4.43	0.0166
product		1	2018.400		2018.400		13.72	0.0005
age*product		2	1384.300	000	692.150	000	4.70	0.0131

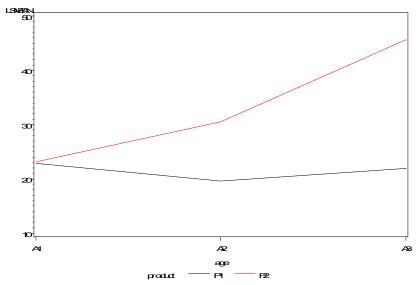
					Standard		
Parameter			Estimate		Error	t Value	Pr > t
Intercept			45.60000000	В	3.83550663	11.89	<.0001
age	A1		-22.50000000	В	5.42422550	-4.15	0.0001
age	A2		-15.10000000	В	5.42422550	-2.78	0.0074
age	A3		0.00000000	В		•	•
product	P1		-23.70000000	В	5.42422550	-4.37	<.0001
product	P2		0.00000000	В		•	•
age*product	A1	P1	23.50000000	В	7.67101326	3.06	0.0034
age*product	A1	P2	0.00000000	В			•
age*product	A2	P1	12.80000000	В	7.67101326	1.67	0.1010
age*product	A2	P2	0.00000000	В			•
age*product	A3	P1	0.00000000	В		•	•
age*product	A3	P2	0.00000000	В		•	•

NOTE: The X'X matrix has been found to be singular, and a generalized inverse was used to solve the normal equations. Terms whose estimates are followed by the letter 'B' are not uniquely estimable.

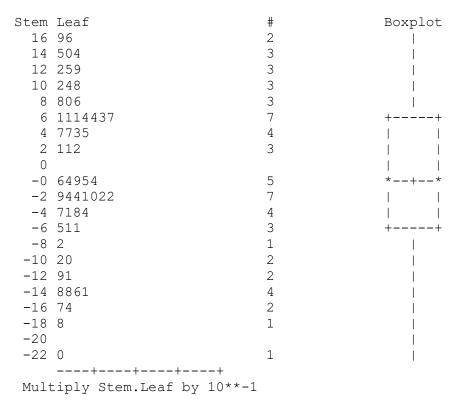
The GLM Procedure Least Squares Means

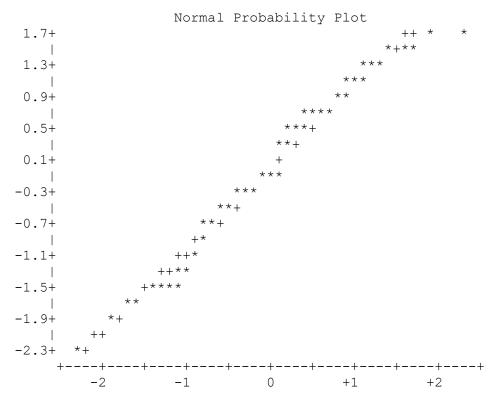
age	product	time LSMEAN
3 1	D.1	00 000000
A1	P1	22.9000000
A1	P2	23.1000000
A2	P1	19.6000000
A2	P2	30.5000000
A3	P1	21.9000000
A3	P2	45.6000000



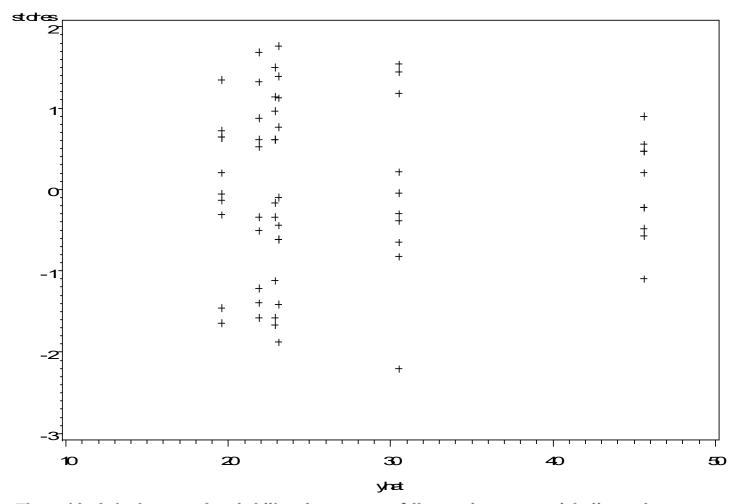


The p-value for the interaction team is 0.013. There is significant evidence of an interaction between the factors Age and Product Type. Thus, the size of the difference between mean attention span of children viewing breakfast cereal ads and viewing video game ads would be different for the three age groups. From the profile plots, the estimated mean attention span for video games is larger than for breakfast cereals, with the size of the difference becoming larger as age increases.





Profile Interaction Prots



The residuals in the normal probability plot appear to fall very close to a straight line and hence we can conclude there is not significant evidence that the residuals have a non-normal distribution.

The plot of the residuals vs. Fitted Value appears to have a consistent width across the fitted values. The condition of constant variance does not appear to be violated.

2X2 in a RCBD – block = piece of metal ... response = block angle type angle*type proc glm data=lathe; class Piece type angle; model y = Piece type|angle; lsmeans type*angle / out=abmeans; lsmeans type / out=ameans; lsmeans angle /out=bmeans; lsmeans type angle / pdiff cl adj=bon; output out=residata p=yhat rstudent=stdres; proc univariate data= residata plots; var stdres; run; proc gplot data=residata; plot stdres*yhat; run; title2 "Profile/Interaction Plots"; symbol1 i=j l=1 v=star c=blue; *draw lines between joint means; symbol2 i=j l=3 v=plus c=red; *draw lines between joint means; proc gplot data=abmeans; plot lsmean*type=angle; plot lsmean*angle=type; run; title2 "Main Effects Plots"; proc gplot data=ameans; plot lsmean*type; run; proc gplot data=bmeans; plot lsmean*angle; run;

E1Q6 Randomized Complete Block With Two Factors

The GLM Procedure

Class Level Information

Class	Levels	Values
Piece	9	1 2 3 4 5 6 7 8 9
type	2	Continuo Interrup
angle	2	15 30

Number of Observations Read 36
Number of Observations Used 36
E1Q6 Randomized Complete Block With Two Factors

The GLM Procedure

Dependent Variable: y

Source		DF	Sum Squa	n of ires	Mean	Square	F Value	Pr > F
Model		11	2111.682	2222	191.	971111	3.07	0.0104
Error		24	1498.447778		62.435324			
Corrected To	tal	35	3610.130	0000				
	R-Square 0.584932	Coeff	T Var	Root 7.901		_	Mean	
	0.304932	20.3	3000	7.901	1003	27.0	00333	
Source		DF	Type I	SS	Mean	Square	F Value	Pr > F
Piece type angle type*angle		8 1 1 1	1510.690 326.404 134.560 140.027	1444	326. 134.	836250 404444 560000 027778	3.02 5.23 2.16 2.24	0.0169 0.0313 0.1551 0.1473
Source		DF	Type III	SS	Mean	Square	F Value	Pr > F
Piece type angle type*angle		8 1 1	1510.690 326.404 134.560 140.027	1444	326. 134.	836250 404444 560000 027778	3.02 5.23 2.16 2.24	0.0169 0.0313 0.1551 0.1473

Significant evidence to suggest a main effect for type of cut \dots

The GLM Procedure Least Squares Means

type	angle	y LSMEAN
Continuo	15	26.9888889
Continuo	30	34.8000000
Interrup	15	24.9111111
Interrup	30	24.8333333

The GLM Procedure Least Squares Means

type	y LSMEAN
Continuo	30.8944444
Interrup	24.8722222

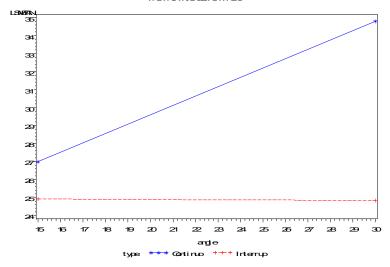
The GLM Procedure Least Squares Means

angle	y LSMEAN
15	25.9500000
30	29.8166667

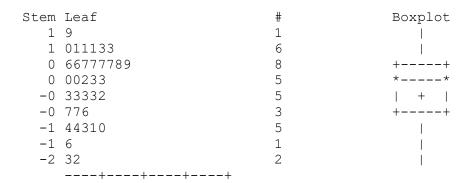
Least Squares Means for Effect type

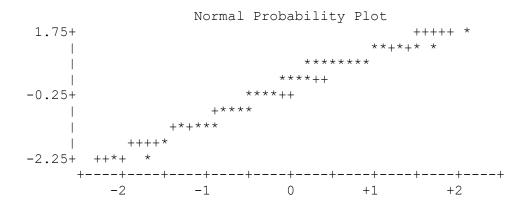
		Difference	Simultane	ous 95%
		Between	Confidence L	imits for
i	j	Means	LSMean(i)-L	SMean(j)
1	2	6.022222	0.586187	11.458258

E1Q6 Randomized Complete Block With Two Factors

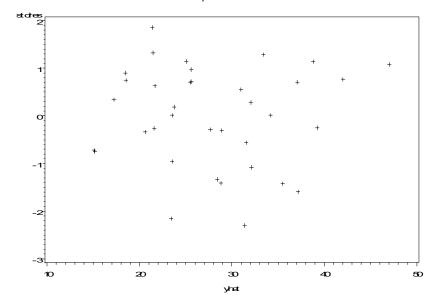


E1Q6 Randomized Complete Block With Two Factors





E1Q6 Randomized Complete Block With Two Factors



Based upon the residual analysis – normality and equal variances for the errors appears reasonable ...

Latin Square to control for two extraneous sources of variation ...

the field must be able to be divided into t^2 plots to handle t treatments concern about too large of plots and too small of plots ...

moisture 5 soil 5 variety 5 error 20 total 35

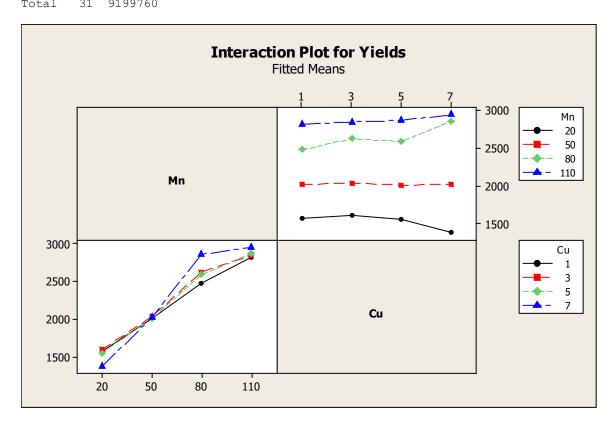
if possible add the control to the experiment so have 7 trts in a 7x7 Latin Square perhaps remove a variety from the experiment to maintain the 6x6 Latin Square ...

14.20

a. The design is a completely randomized 4x4 factorial experiment with Factor A-Cu Rate and Factor B- Mn Rate. There are two replications of the 16 treatments.

Analysis of Variance for Yields, using Adjusted SS for Tests

Source	DF	Seq SS	Adj SS	Adj MS	F	P
Mn	3	8935108	8935108	2978369	1486.70	0.000
Cu	3	28199	28199	9400	4.69	0.016
Mn*Cu	9	204399	204399	22711	11.34	0.000
Error	16	32053	32053	2003		
Total	21	0100760				



Based on the profile plot, there appears to be a strong interaction between the factors Cu Rate and Mn Rate. The mean soybean yield increases for increasing Cu Rate at a Mn Rate of 80 but the mean soybean yield stays constant initially and then decreases for increasing Cu Rate at a Mn Rate of 20. At a Mn Rate of 110 and 50, the mean soybean yield remains relatively constant with increasing rates of Cu.

14.21

a. The test for interaction yields p-value 0.0001. This implies there is significant evidence of an interaction between Cu Rate and Mn Rate on Soybean yield.

b.
$$Mn = 110$$

$$c. Cu = 7$$

d.
$$(Cu,Mn) = (7,110)$$

Evaluation #1 SKETCH OF SOULTIONS 9)

2-way random effects model

MPG = overall mean + driver + car + driver*car + error

driver, car, driver*car = random effects

E (MS)

Source Type III Expected Mean Square

driver Var(Error) + 2 Var(driver*car) + 10 Var(driver)

car Var(Error) + 2 Var(driver*car) + 8 Var(car)

driver*car Var(Error) + 2 Var(driver*car)

Dependent Variable: mpg

Sum of

Source DF Squares Mean Square F Value Pr > F

Model 19 377.4447500 19.8655132 113.03 <.0001

Error 20 3.5150000 0.1757500

Corrected Total 39 380.9597500

The GLM Procedure

Tests of Hypotheses for Random Model Analysis of Variance

Dependent Variable: mpg

Source DF Type III SS Mean Square F Value Pr > F

driver 3 280.284750 93.428250 458.26 <.0001 car 4 94.713500 23.678375 116.14 <.0001

Error 12 2.446500 0.203875

Error: MS(driver*car)

Source DF Type III SS Mean Square F Value Pr > F

driver*car 12 2.446500 0.203875 1.16 0.3715

Error: MS(Error) 20 3.515000 0.175750

The Mixed Procedure

Covariance Parameter Estimates

Cov Parm Estimate Alpha Lower Upper

 driver
 9.3224
 0.05
 2.9864
 130.79

 car
 2.9343
 0.05
 1.0464
 24.9038

 driver*car
 0.01406
 0.05
 0.001345
 3.592E17

 Residual
 0.1757
 0.05
 0.1029
 0.3665

2-way random effects model

```
Main Effects Variance Components - Significant
Interaction Variance Component - Non-Significant
Driver Variance Component - greater effect ...

proc glm data = mpg;
class driver car;
model mpg = driver car driver*car;
random driver car driver*car / test;
run;

proc mixed data = mpg cl;
class driver car;
model mpg =;
random driver car driver*car;
run; quit;
```